NOTES ON PRINCIPAL COMPONENT ANALYSIS

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1. NOTATION

We use Einstein notation [\(link\)](https://en.wikipedia.org/wiki/Einstein_notation) for vectors and matrices. Specifically, suppose **v** is a vector in **R***N*. Unless specified otherwise, we consider **v** to be a column vector with entries indicated by upper indices:

$$
\mathbf{v}_i = \begin{bmatrix} v^1 \\ v^2 \\ \vdots \\ v^N \end{bmatrix}
$$

It will be clear from context whether the upper subscripts denote indices or exponents. To save space, we may also write **v** as a tuple: $\mathbf{v} = (v^1, v^2, \dots, v^N)$. The mean of the vector **v** is defined as:

$$
\text{mean}(\mathbf{v}) = \frac{1}{N} \sum_{i=1}^{N} v^i
$$

The covariance of two vectors **v** and **w** in \mathbb{R}^N is defined as:

$$
cov(\mathbf{v}, \mathbf{w}) = \frac{1}{N} \sum_{i=1}^{N} (v^i - \text{mean}(\mathbf{v})) (w^i - \text{mean}(\mathbf{w}))
$$

2. The data matrix *X*

Let $X \in \mathbb{R}^{N \times d}$ be the data matrix. There are *N* samples, recorded as the rows of *X*, and each sample has *d* features, corresponding to the columns of *X*. We assume $N \ge d$. Following Einstein notation we denote matrix entries using upper indices for the rows and lower indices for the columns:

$$
X = \begin{bmatrix} x_1^1 & x_2^1 & x_3^1 & \cdots & x_d^1 \\ x_1^2 & x_2^2 & x_3^2 & \cdots & x_d^1 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ x_1^N & x_2^N & x_3^N & \cdots & x_d^N \end{bmatrix} \in \mathbb{R}^{N \times d}
$$

In other words, x_i^i *j* denotes the (scalar) entry appearing in the *i*-th column and *j*-th row of *X*, for $i = 1, \ldots, N$ and $j = 1, \ldots, d$. This is the measurement of the *d*-th feature of the *i*-th sample. The covariance matrix^{[1](#page-0-0)} of *X* is a *d* by *d* matrix whose (j, k) entry is given by

¹The covariance matrix can be computed in numpy via the command np.cov(X, ddof=0). Having delta degrees of freedom (ddof) equal to *d* amounts to dividing by *N* − *d* instead of *N*. The default value is ddof $= 1$, which gives an unbiased estimator for the population covariance.

the covariance of the *j*-th and *k*-th columns of *X*:

$$
Covmat(X)kj = cov ((xj1,..., xjN),(xk1,..., xkN))
$$

3. The matrix *A*

Let $Id_{N \times N}$ be the identity *N* by *N* matrix, and let $1\!\!1_{N \times N}$ be the *N* by *N* matrix of all ones. Define a new *N* by *N* matrix as:

$$
A = \mathrm{Id}_{N \times N} - \frac{1}{N} \mathbb{1}_{N \times N}
$$

We leave the proof of the following lemma as an exercise:

Lemma 3.1. *We have:*

- (1) *The matrix A is symmetric, and* $A^2 = A$.
- (2) *Each column of the matrix AX has mean zero.*
- (3) *The covariance of two vectors* **v** *and* **w** *in* \mathbb{R}^N *is given by:*

$$
cov(\mathbf{v}, \mathbf{w}) = \frac{1}{N} \mathbf{v}^T A \mathbf{w}
$$

(4) *The covariance matrix of X is given by:*

$$
Covmat(X) = \frac{1}{N} X^T A X
$$

4. Principal Component Analysis

Lemma 3.[1](#page-1-0).2, shows that the matrix *AX* is 'centered', i.e., its column means are all zero. Let $AX = USV^T$ be the singular value decomposition of AX. Hence, $U \in \mathbb{R}^{N \times N}$ is an orthogonal matrix, *S* ∈ **R***N*×*^d* is a diagonal matrix with non-negative, non-increasing entries along the diagonal (the singular values σ_i of AX), and $V \in \mathbb{R}^{d \times d}$ is an orthogonal matrix. Terminology:

• The columns of *V* are the *pincipal components* of *X*.

Using Lemma [3](#page-1-0).1.4, and the fact that $U^TU = Id_{N \times N}$, one computes that the covariance matrix of *AXV* is given by:

$$
Covmat(AXV) = \frac{S^TS}{N} \in \mathbb{R}^{d \times d}
$$

Since *S* is a diagonal matrix, we conclude that Covmat(*AXV*) is also a diagonal matrix. Its *i*-th diagonal entry, denoted *λⁱ* , is the square of the *i*-th singular value of *AX*, divided by *N*, that is, $\lambda_i = \frac{\sigma_i^2}{N}$ for $i = 1, \ldots, d$. Terminology:

• The dataset $R = AXY$ is the *centered, diagonalized* version of *X*.

Assume all singular values of *X* are positive (this is usually the case in practice). Then let $\Lambda^{-1/2}$ be the diagonal *d* by *d* matrix whose *i*-th entry is $(\lambda_i)^{-1/2} = \frac{\sqrt{N}}{\sigma_i}$ *σi* . We compute that the covariance matrix of $AXV\Lambda^{-1/2}$ $AXV\Lambda^{-1/2}$ $AXV\Lambda^{-1/2}$ is the identity matrix²:

$$
Covmat(AXV\Lambda^{-1/2}) = Id_{d \times d}
$$

Terminology:

• The dataset $Z = AXV\Lambda^{-1/2}$ is the *whitened* version of *X*, with mean zero and unit variance.

5. Projections

Let $s \leq d$, and consider the matrix:

$$
\pi_s = \begin{bmatrix} \mathrm{Id}_{s\times s} & 0 \\ 0 & 0 \end{bmatrix}
$$

This is a projection matrix onto the first *s* components. We now:

- Project the centered, diagonalized data to obtain *AXVπ^s* .
- Apply V^T to obtain the projected centered, undiagonalized data: $AXV\pi_sV^T$.
- Add the column means to obtain the *s-truncated* version of *X*:

$$
\hat{X} = AXV\pi_s V^T + \frac{1}{N}1\!\!1_{N \times N}X
$$

Note that, if $s = d$, then $\hat{X} = X$, and if $s = 0$, then \hat{X} is the matrix of column means. Otherwise, \hat{X} is a lower-dimensional representation of *X*. One can show that \hat{X} simplifies to:

$$
\hat{X} = X - US\left(\mathrm{Id}_{d \times d} - \pi_s\right) V^T.
$$

In other words, one zeros out the first *s* singular values of $AX = USV^T$ and subtract the result from *X*.

Finally, we compute the error of this lower-dimensional representation. Let $\mathbf{x}^{(i)}$ denote the *i*-th row of *X*, representing the *i* sample. Similarly, let $\hat{\mathbf{x}}^{(i)}$ denote the *i*-th row of \hat{X} . Then the error is:

$$
\begin{split} \text{Error} &= \frac{1}{N} \sum_{i=1}^{N} \|\mathbf{x}^{(i)} - \hat{\mathbf{x}}^{(i)}\|^2 = \frac{1}{N} \sum_{i=1}^{N} (\mathbf{x}^{(i)} - \hat{\mathbf{x}}^{(i)})^T (\mathbf{x}^{(i)} - \hat{\mathbf{x}}^{(i)}) \\ &= \frac{1}{N} \text{Trace}\left(\left(X - \hat{X} \right)^T (X - \hat{X}) \right) \\ &= \frac{1}{N} \text{Trace}\left(\left(\text{US} \left(\text{Id}_{d \times d} - \pi_s \right) V^T \right)^T \left(\text{US} \left(\text{Id}_{d \times d} - \pi_s \right) V^T \right) \right) \\ &= \frac{1}{N} \text{Trace}\left(S^T S - S^T S \pi_s \right) = \sum_{i=s+1}^d \lambda_i \end{split}
$$

²The same is true for *AXV*Λ−1/2*^Q* where *^Q* [∈] **^R***d*×*^d* is any orthogonal matrix